

# **Biips**: A software for Bayesian inference with interacting particle systems Probabilistic Programming Reading Group

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## Outline

### Context

### Graphical models and BUGS language

## SMC

### Matbiips

### Particle MCMC

# Summary

#### Context

Graphical models and BUGS language

### SMC

Matbiips

Particle MCMC

## Context

**Biips** = Bayesian inference with interacting particle systems

### Bayesian inference

- Sample from a posterior distribution  $p(X|Y) = \frac{p(X,Y)}{p(Y)}$
- High dimensional, arbitrary complexity
- Simulation methods: MCMC, SMC...

### Motivation

- Last 20 years: success of SMC in many applications
- ▶ No general and easy-to-use software for SMC

## Context

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### Bayesian inference

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### Motivation

- Last 20 years: success of SMC in many applications
- No general and easy-to-use software for SMC

## Context

### **Biips** = Bayesian inference with interacting particle systems

### Objectives

- BUGS language compatible
- Extensibility: user-defined functions/samplers
- Black-box SMC inference engine
- Interfaces with popular software: Matlab/Octave, R
- Post-processing

# Summary

#### Context

### Graphical models and BUGS language

SMC

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Particle MCMC



The graph displays a factorization of the joint distribution:

 $p(x_{1:3},y_{1:2})$ 

Directed acyclic graph



Directed acyclic graph

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 $p(x_{1:3}, y_{1:2}) = p(x_1) \ p(x_2|x_1) \ p(y_1|x_2) \ p(x_3|x_1, x_2) \ p(y_2|x_2, x_3)$ 



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- Stochastic relations
- Deterministic relations

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```
Linear regression:
model {
    Y ~ dnorm(mu, tau)
```



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```
Linear regression:
model {
    Y ~ dnorm(mu, tau)
    tau ~ dgamma(0.01, 0.01)
```



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```
Linear regression:
model {
    Y ~ dnorm(mu, tau)
    tau ~ dgamma(0.01, 0.01)
    mu <- beta * X + alpha</pre>
```



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```
Linear regression:
model {
    Y ~ dnorm(mu, tau)
    tau ~ dgamma(0.01, 0.01)
    mu <- beta * X + alpha
    alpha ~ dnorm(0, 1E-6)
    beta ~ dnorm(0, 1E-6)
```



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```
Linear regression:
model {
    Y ~ dnorm(mu, tau)
    tau ~ dgamma(0.01, 0.01)
    mu <- beta * X + alpha
    alpha ~ dnorm(0, 1E-6)
    beta ~ dnorm(0, 1E-6)
}</pre>
```

```
Goal:
Estimate p(\alpha, \beta, \tau | X, Y)
```



1E-6

0

X

 $\beta X$ 

BUGS software using MCMC

### **BUGS** = **B**ayesian inference **U**sing **G**ibbs **S**ampling

- WinBUGS, OpenBUGS, JAGS [Plummer, 2012]
- Expert system automatically derives MCMC methods (Gibbs, Slice, Metropolis, ...) in a 'black-box' fashion
- Very popular among practitioners, applying MCMC methods to a wide range of applications [Lunn et al., 2012]

# Summary

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### SMC

#### Matbiips

#### Particle MCMC

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Topological sort (with priority to measurement nodes):  $(X_1, Y_1, Y_3, X_3, X_2, Y_4, Y_2)$ 



Rearrangement of the directed acyclic graph:



Topological sort (with priority to measurement nodes):  $\underbrace{(X_1, \underbrace{Y_1, Y_3}_{\mathbf{X'_1}}, \underbrace{X_3, X_2}_{\mathbf{X'_2}}, \underbrace{Y_4, Y_2}_{\mathbf{Y'_2}})}_{\mathbf{X'_1}}$ 



Topological sort (with priority to measurement nodes):  $(\underbrace{X_1}_{\mathbf{X}_1}, \underbrace{Y_1}_{\mathbf{Y}_1}, \underbrace{X_3}_{\mathbf{X}_2}, \underbrace{X_2}_{\mathbf{X}_2}, \underbrace{Y_4}_{\mathbf{Y}_2}, \underbrace{Y_2}_{\mathbf{Y}_2})$  Rearrangement of the directed acyclic graph:



The statistical model decomposes as  $p(x'_1, x'_2, y'_1, y'_2) =$  $p(x'_1)p(y'_1|x'_1)$  $p(x'_2|x'_1, y'_1)p(y'_2|x'_2)$ 

## SMC algorithm

More generally, assume that we have sorted variables  $(X_1, Y_1, \ldots, X_n, Y_n)$ . The statistical model decomposes as

$$p(x_{1:n},y_{1:n})=p(x_1)p(y_1|x_1)\prod_{t=2}^n p(x_t|\mathsf{pa}(x_t))p(y_t|\mathsf{pa}(y_t))$$

where pa(x) denotes the set of parents of variable x.

# SMC algorithm

- A.k.a. interacting MCMC, particle filtering, sequential Monte Carlo methods (SMC) ...
- Sequentially sample from conditional distributions of increasing dimension

 $\pi_1(x_1|y_1) o \pi_2(x_{1:2}|y_{1:2}) o ... o \pi_n(x_{1:n}|y_{1:n})$ 

where, for t = 1, ..., n

$$\pi_t(x_{1:t}|y_{1:t}) = rac{p(x_{1:t},y_{1:t})}{p(y_{1:t})}$$

Two stochastic mechanisms:

- Mutation/Exploration
- Selection

[Doucet et al., 2001, Del Moral, 2004, Doucet and Johansen, 2010]  $$^{13/41}$$ 

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# SMC Algorithm

## Standard SMC algorithm

For  $t = 1, \ldots, n$ 

▶ For i = 1,..., N
▶ Sample: X<sup>(i)</sup><sub>t,t</sub> ~ q<sub>t</sub> and let X<sup>(i)</sup><sub>t,1:t</sub> = (X<sup>(i)</sup><sub>t-1,1:t-1</sub>, X<sup>(i)</sup><sub>t,t</sub>)
▶ Weight: w<sup>(i)</sup><sub>t</sub> = π(y<sub>t</sub>|pa(y<sub>t</sub>))π(x<sup>(i)</sup><sub>t,t</sub>|pa(x<sup>(i)</sup><sub>t,t</sub>)))/q<sub>t</sub>(x<sup>(i)</sup><sub>t,t</sub>)
▶ Normalize: W<sup>(i)</sup><sub>t</sub> = w<sup>(i)</sup><sub>∑<sub>j=1</sub><sup>N</sup> w<sup>(j)</sup><sub>t</sub>
▶ Resample: {X<sup>(i)</sup><sub>t,1:t</sub>, W<sup>(i)</sup><sub>t</sub>} i=1,...,N → {X<sup>(i)</sup><sub>t,1:t</sub>, 1/N} i=1,...,N
</sub>

### Outputs

- Weighted particles  $(W_t^{(i)}, X_{t,1:t}^{(i)})_{i=1,...,N}$  for  $t=1,\ldots,n$
- Estimate of the marginal likelihood  $\widehat{Z} = \prod_{t=1}^n \left( rac{1}{N} \sum_{i=1}^N w_t^{(i)} 
  ight)$

# SMC algorithm

### Marginal distributions

$$\pi_1(x_1|y_1) \ o \ \pi_2(x_{1:2}|y_{1:2}) \ o ... o \ \pi_n(x_{1:n}|y_{1:n})$$








































## Limitations and diagnosis of SMC algorithms



For a given  $t \leq n$ , for each unique value  $X_{n,t}^{\prime(k)}$ ,  $k = 1, \ldots, K_{n,t}$ , let  $W_{n,t}^{\prime(k)} = \sum_{i|X_t^{(i)}=X_t^{\prime(k)}} W_n^{(i)}$  be its associated total weight. A measure of the quality of the approximation of the posterior distribution  $p(x_{t:n}|y_{1:n})$  is given by the smoothing effective sample size (SESS):

$$SESS_{t} = \frac{1}{\sum_{k=1}^{K_{n,t}} \left( W_{n,t}^{\prime(k)} \right)^{2}}$$
(1)

with  $1 \leq \text{SESS}_t \leq N$ .

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## Technical implementation

![](_page_51_Figure_1.jpeg)

- Interfaces: Matlab/Octave, R
- Multi-platform: Windows, Linux, Mac OSX
- Free and open source (GPL)

### Switching Stochastic Volatility (SSV)

Let  $Y_t$  be the response variable and  $X_t$  the unobserved log-volatility of  $Y_t$ . For  $t=1,\ldots,n$ 

$$egin{aligned} X_t | (X_{t-1} = x_{t-1}, C_t = c_t) &\sim \mathcal{N}(lpha_{c_t} + \phi x_{t-1}, \sigma^2) \ Y_t | X_t = x_t &\sim \mathcal{N}(0, \exp(x_t)) \end{aligned}$$

The regime variables  $C_t$  follow a two-state Markov process with transition probabilities

 $p_{ij} = \Pr(C_t = j | C_{t-1} = i), \text{ for } i, j = 1, 2$ 

![](_page_52_Figure_5.jpeg)

## SSV model in BUGS language

switch\_stoch\_volatility.bug

```
model
{
    c[1] ~ dcat(pi[c0,])
    mu[1] <- alpha[1]*(c[1]==1) + alpha[2]*(c[1]==2) + phi*x0
    x[1] ~ dnorm(mu[1], 1/sigma^2)
    y[1] ~ dnorm(0, exp(-x[1]))
    for (t in 2:t_max)
    {
        c[t] ~ dcat(ifelse(c[t-1]==1, pi[1,], pi[2,]))
        mu[t] <- alpha[1]*(c[t]==1) + alpha[2]*(c[t]==2) + phi*x[t-1]
        x[t] ~ dnorm(mu[t], 1/sigma^2)
        y[t] ~ dnorm(0, exp(-x[t]))
    }
}</pre>
```

### Model compilation

```
Matbiips

sigma = .4; alpha = [-2.5; -1]; phi = .5; c0 = 1; x0 = 0; t_max =

200;

pi = [.9, .1; .1, .9];

data = struct('t_max', t_max, 'sigma', sigma,...

'alpha', alpha, 'phi', phi, 'pi', pi, 'c0', c0, 'x0', x0);

model_file = 'switch_stoch_volatility.bug';

model = biips_model(model_file, data, 'sample_data', true);

data = model.data:
```

## SMC samples

Matbiips

```
n_part = 5000;
variables = {'x'};
out_smc = biips_smc_samples(model, variables, n_part);
diag_smc = biips_diagnosis(out_smc);
```

![](_page_55_Figure_3.jpeg)

#### Summary statistics

summ = biips\_summary(out\_smc, 'probs', [.025, .975]); x\_f\_mean = summ.x.f.mean; x\_f\_quant = summ.x.f.quant; x\_s\_mean = summ.x.s.mean; x\_s\_quant = summ.x.s.quant;

![](_page_56_Figure_2.jpeg)

Matbiips

### Kernel density estimates

![](_page_57_Figure_1.jpeg)

### Sensitivity analysis

```
Matbiips
```

![](_page_58_Figure_3.jpeg)

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## Summary

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Matbiips

Particle MCMC

Recent algorithms that use SMC algorithms within a MCMC algorithm

- Particle Independant Metropolis-Hastings (PIMH)
- Particle Marginal Metropolis-Hastings (PMMH)

#### Static parameter estimation

![](_page_61_Figure_1.jpeg)

Due to the successive resamplings, SMC estimations of  $p(\theta|y_{1:n})$  might be poor.

The PMMH splits the variables in the graphical model into two sets:

- $\blacktriangleright$  a set of variables X that will be sampled using a SMC algorithm
- a set  $\theta = (\theta_1, \dots, \theta_p)$  sampled with a MH proposal

## PMMH

Standard PMMH algorithm Set  $\widehat{Z}(0) = 0$  and initialize  $\theta(0)$ For  $k = 1, \dots, n_{\text{iter}}$ ,

- Sample  $\theta^{\star} \sim \nu$
- ▶ Run a SMC to approximate  $p(x_{1:n}|y_{1:n}, \theta^{\star})$  with output  $(X_{1:n}^{\star(i)}, W_n^{\star(i)})_{i=1,...,N}$  and  $\widehat{Z}^{\star}$
- With probability

$$\min\left(1,rac{\widehat{Z}^{\star}}{\widehat{Z}(k-1)}
ight)$$

set  $X_{1:n}(k) = X_{1:n}^{\star(\ell)}$ ,  $\theta(k) = \theta^{\star}$  and  $\widehat{Z}(k-1) = \widehat{Z}^{\star}$ , where  $\ell \sim \operatorname{Discrete}(W_n^{\star(1)}, \dots, W_n^{\star(N)})$ 

otherwise, keep previous iteration values

#### Outputs

• MCMC samples  $(X_{1:n}(k), \theta(k))_{k=1,...,n_{\mathsf{iter}}}$ 

#### Static parameter estimation in the SSV model

We consider the following prior on the parameters lpha,  $\pi$ ,  $\phi$  and au :

 $egin{aligned} lpha_1 &= \gamma_1 \ lpha_2 &= \gamma_1 + \gamma_2 \ \gamma_1 &\sim \mathcal{N}(0, 100) \ \gamma_2 &\sim \mathcal{TN}_{(0, +\infty)}(0, 100) \end{aligned}$ 

 $\begin{aligned} \frac{1}{\sigma^2} &\sim \text{Gamma}(2.001, 1) \\ \phi &\sim \mathcal{TN}_{(-1,1)}(0, 100) \\ \pi_{11} &\sim \text{Beta}(10, .5) \\ \pi_{22} &\sim \text{Beta}(10, .5) \end{aligned}$ 

## SSV model with unknown parameters in BUGS language

switch\_stoch\_volatility\_param.bug

```
model
ł
  gamma[1] ~ dnorm(0, 1/100)
  gamma[2] ~ dnorm(0, 1/100) T(0,)
  alpha[1] <- gamma[1]
  alpha[2] <- gamma[1] + gamma[2]</pre>
  phi ~ dnorm(0, 1/100) T(-1,1)
  tau ~ dgamma(2.001, 1)
  sigma <- 1/sqrt(tau)</pre>
  pi[1,1] ~ dbeta(10, .5)
  pi[1,2] <- 1.00 - pi[1,1]
  pi[2,2] ~ dbeta(10, .5)
  pi[2,1] <- 1.00 - pi[2,2]
  . . .
```

```
Matbiips
```

```
model_file = 'switch_stoch_volatility_param.bug';
model = biips_model(model_file, data, 'sample_data', sample_data);
data = model.data;
```

### **PMMH** samples

```
Run a PMMH sampler to approximate p(\alpha_1, \alpha_2, \sigma, \pi_{11}, \pi_{22}, \phi, X_{1,T}, C_{1:T} | Y_{1:T}).
```

```
Matbiips
```

```
n burn = 2000;
n iter = 40000;
thin = 10:
n part = 50;
param_names = {'gamma[1,1]', 'gamma[2,1]', 'phi', 'tau', 'pi[1,1]',
    'pi[2,2]'};
latent_names = {'x', 'alpha[1,1]', 'alpha[2,1]', 'sigma'};
inits = \{-1, 1, .5, 5, .8, .8\}:
obj_pmmh = biips_pmmh_init(model, param_names, 'inits', inits, '
    latent names'. latent names):
obj_pmmh = biips_pmmh_update(obj_pmmh, n_burn, n_part);
[obj_pmmh, out_pmmh, log_marg_like_pen, log_marg_like] = ...
    biips pmmh samples (obj pmmh, n iter, n part, 'thin', thin);
```

#### Posterior samples

![](_page_66_Figure_1.jpeg)

## Other features of Biips

- Backward smoothing algorithm
- Particle Independent Metropolis-Hastings algorithm
- Automatic choice of the proposal distribution including
   Optimal/Conditional samplers: Gaussian-Gaussian, Beta-Bernoulli, Finite discrete
- ► Easy BUGS language extensions with user-defined Matlab/R functions

## Related software

#### using MCMC

- WinBUGS, OpenBUGS [Lunn et al., 2000, Lunn et al., 2012], JAGS [Plummer, 2003]
- Stan [Stan Development Team, 2013]

#### using SMC

- SMCTC [Johansen, 2009]
- LibBi [Murray, 2013]

#### using both

▶ Venture [Mansinghka et al., 2014], Anglican [Wood et al., 2014]

## Conclusion

- BUGS language compatible
- Extensibility: user-defined functions/samplers
- Black-box SMC inference engine
- ► Interfaces with popular software: Matlab/Octave, R
- Post-processing

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![](_page_70_Picture_3.jpeg)

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![](_page_70_Picture_5.jpeg)

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![](_page_70_Picture_8.jpeg)

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![](_page_70_Picture_10.jpeg)

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![](_page_71_Picture_5.jpeg)

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![](_page_71_Picture_8.jpeg)

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![](_page_71_Picture_11.jpeg)

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## THANK YOU



http://alea.bordeaux.inria.fr/biips