



Biips software: **B**ayesian **i**nference with **i**nteracting
particle **s**ystems

BAYES 2015

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Basel, May 2015

Outline

Context

Graphical models and BUGS language

SMC

Biips software

Particle MCMC

Summary

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Graphical models and BUGS language

SMC

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Context

Biips = **B**ayesian **i**nference with **i**nteracting **p**article **s**ystems

Bayesian inference

- ▶ Sample from a posterior distribution $p(\mathbf{X} | \mathbf{Y}) = \frac{p(\mathbf{X}, \mathbf{Y})}{p(\mathbf{Y})} = \frac{p(\mathbf{X}, \mathbf{Y})}{Z}$
- ▶ High dimensional, arbitrary complexity
- ▶ Simulation methods: MCMC, SMC...

Motivation

- ▶ Last 20 years: success of SMC in many applications
- ▶ No general and easy-to-use software for SMC

Context

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Features

- ▶ BUGS language compatible
- ▶ Extensibility: custom functions/samplers
- ▶ Black-box SMC inference engine
- ▶ Interfaces with popular software: Matlab/Octave, R
- ▶ Post-processing tools

Summary

Context

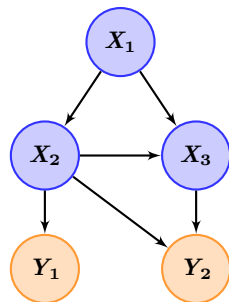
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Graphical models

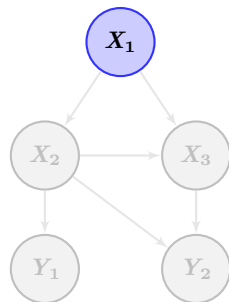


Directed acyclic graph

The graph displays a **factorization** of the joint distribution:

$$p(x_{1:3}, y_{1:2})$$

Graphical models

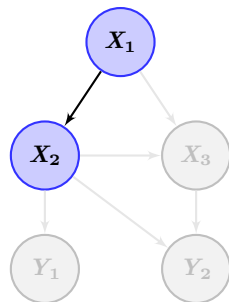


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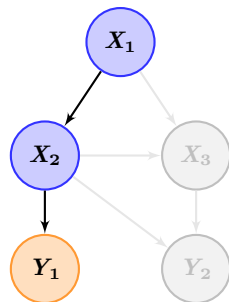


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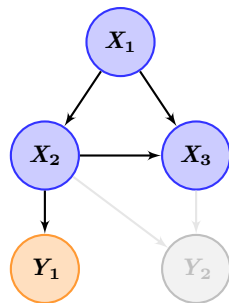


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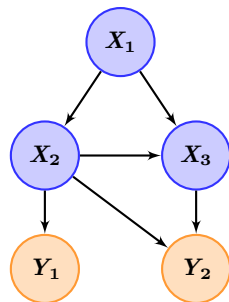


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BUGS language

- ▶ S-like declarative language for describing graphical models
- ▶ Stochastic relations
- ▶ Deterministic relations

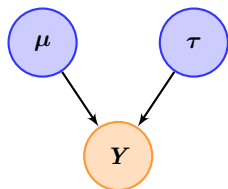
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Linear regression:

```
model {  
    Y ~ dnorm(mu, tau)
```

```
}
```

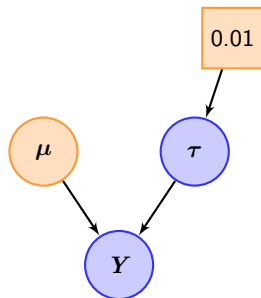


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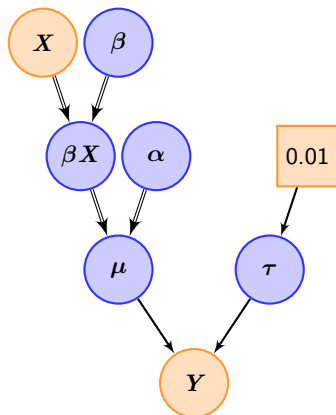


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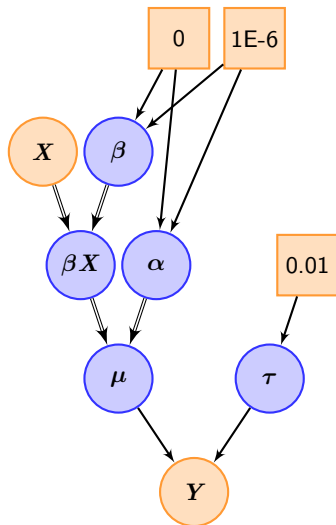


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BUGS language

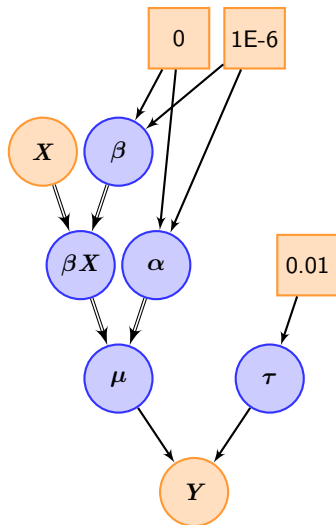
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}
```

Goal:

Estimate $p(\alpha, \beta, \tau | X, Y)$



BUGS software using MCMC

BUGS = **B**ayesian inference **U**sing **G**ibbs **S**ampling

- ▶ WinBUGS, OpenBUGS, JAGS [Plummer, 2012]
- ▶ Expert system automatically derives **MCMC methods** (Gibbs, Slice, Metropolis, ...) in a '**black-box**' fashion
- ▶ Very **popular** among practitioners, applying MCMC methods to a wide range of applications [Lunn et al., 2012]

Summary

Context

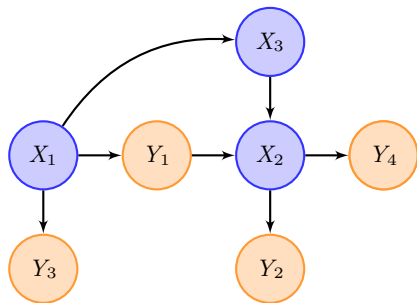
Graphical models and BUGS language

SMC

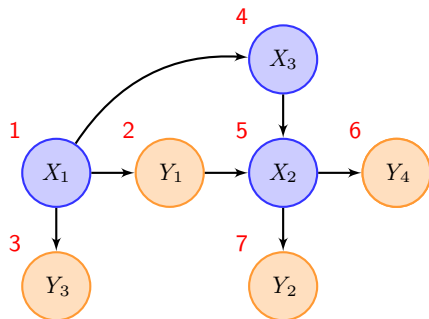
Biips software

Particle MCMC

Ordering of the graph



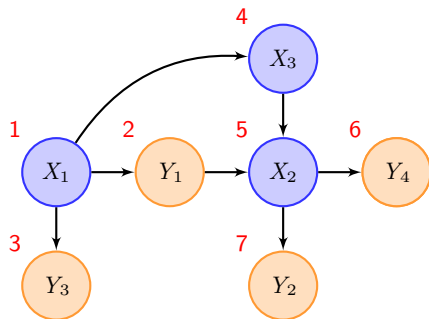
Ordering of the graph



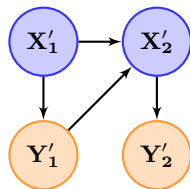
Topological sort (with priority to measurement nodes):

($X_1, Y_1, Y_3, X_3, X_2, Y_4, Y_2$)

Ordering of the graph



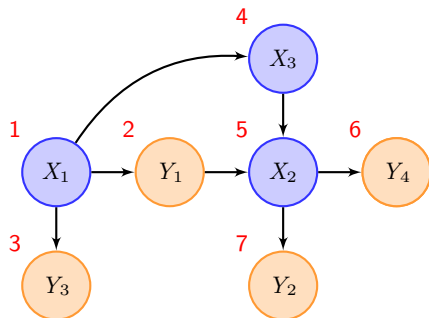
Rearrangement of the directed acyclic graph:



Topological sort (with priority to measurement nodes):

$$\left(\underbrace{X_1}_{X'_1}, \underbrace{Y_1, Y_3}_{Y'_1}, \underbrace{X_3, X_2}_{X'_2}, \underbrace{Y_4, Y_2}_{Y'_2} \right)$$

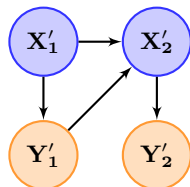
Ordering of the graph



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$$\underbrace{(X_1)}_{X'_1}, \underbrace{(Y_1, Y_3)}_{Y'_1}, \underbrace{(X_3, X_2)}_{X'_2}, \underbrace{(Y_4, Y_2)}_{Y'_2}$$

Rearrangement of the directed acyclic graph:



The statistical model decomposes as

$$\begin{aligned} p(x'_1, x'_2, y'_1, y'_2) &= \\ p(x'_1) &p(y'_1|x'_1) \\ p(x'_2|x'_1, y'_1) &p(y'_2|x'_2) \end{aligned}$$

SMC algorithm

More generally, assume that we have sorted variables $(X_1, Y_1, \dots, X_n, Y_n)$.

The statistical model decomposes as

$$p(\mathbf{x}_{1:n}, \mathbf{y}_{1:n}) = p(\mathbf{x}_1)p(\mathbf{y}_1|\mathbf{x}_1) \prod_{t=2}^n p(\mathbf{x}_t|\text{pa}(\mathbf{x}_t))p(\mathbf{y}_t|\text{pa}(\mathbf{y}_t))$$

where $\text{pa}(\mathbf{x})$ denotes the set of parents of variable \mathbf{x} .

SMC algorithm

- ▶ A.k.a. interacting MCMC, particle filtering, sequential Monte Carlo methods (SMC) ...
- ▶ Sequentially sample from conditional distributions of increasing dimension

$$\pi_1(x_1|y_1) \rightarrow \pi_2(x_{1:2}|y_{1:2}) \rightarrow \dots \rightarrow \pi_n(x_{1:n}|y_{1:n})$$

where, for $t = 1, \dots, n$

$$\begin{aligned}\pi_t(x_{1:t}|y_{1:t}) &= \frac{p(x_{1:t}, y_{1:t})}{p(y_{1:t})} \\ &= \pi_{t-1}(x_{1:t-1}|y_{1:t-1}) \frac{p(x_t|\text{pa}(x_t))p(y_t|\text{pa}(y_t))}{p(y_t|y_{1:t-1})}\end{aligned}$$

Two stochastic mechanisms:

- ▶ **Mutation/Exploration**
- ▶ **Selection**

[Doucet et al., 2001, Del Moral, 2004, Doucet and Johansen, 2010]

Standard SMC Algorithm

For $t = 1, \dots, n$

▶ For $i = 1, \dots, N$

▶ Sample: $X_{t,t}^{(i)} \sim q_t$ and let $X_{t,1:t}^{(i)} = (\tilde{X}_{t-1,1:t-1}^{(i)}, X_{t,t}^{(i)})$

▶ Weight: $w_t^{(i)} = \frac{\pi(y_t | \text{pa}(y_t)) \pi(x_{t,t}^{(i)} | \text{pa}(x_{t,t}^{(i)}))}{q_t(x_{t,t}^{(i)})}$

▶ Normalize: $W_t^{(i)} = \frac{w_t^{(i)}}{\sum_{j=1}^N w_t^{(j)}}$

▶ Resample: $\{X_{t,1:t}^{(i)}, W_t^{(i)}\}_{i=1, \dots, N} \rightarrow \{\tilde{X}_{t,1:t}^{(i)}, \frac{1}{N}\}_{i=1, \dots, N}$

Outputs

▶ Weighted particles $(W_t^{(i)}, X_{t,1:t}^{(i)})_{i=1, \dots, N}$ for $t = 1, \dots, n$

▶ Estimate of the marginal likelihood $\hat{Z} = \prod_{t=1}^n \left(\frac{1}{N} \sum_{i=1}^N w_t^{(i)} \right)$

SMC algorithm

Marginal distributions

$$\pi_1(\mathbf{x}_1|\mathbf{y}_1) \rightarrow \pi_2(\mathbf{x}_{1:2}|\mathbf{y}_{1:2}) \rightarrow \dots \rightarrow \pi_n(\mathbf{x}_{1:n}|\mathbf{y}_{1:n})$$

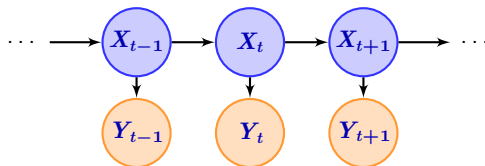
Filtering: $\pi_1(\mathbf{x}_1|\mathbf{y}_1), \quad \pi_2(\mathbf{x}_2|\mathbf{y}_{1:2}), \quad \dots, \quad \pi_n(\mathbf{x}_n|\mathbf{y}_{1:n})$

Smoothing: $\pi_1(\mathbf{x}_1|\mathbf{y}_{1:n}), \quad \pi_2(\mathbf{x}_2|\mathbf{y}_{1:n}), \quad \dots, \quad \pi_n(\mathbf{x}_n|\mathbf{y}_{1:n})$

Example

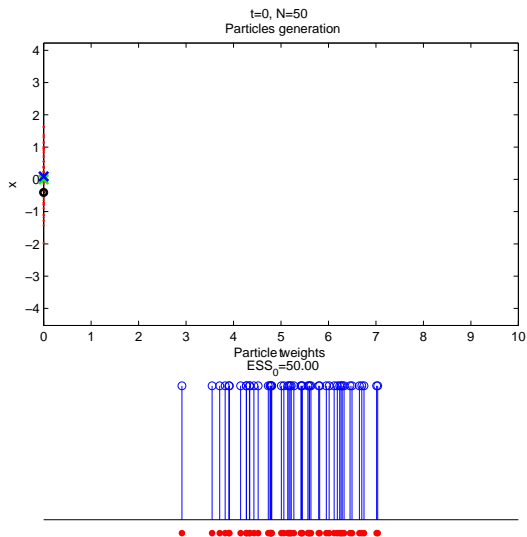
Hidden Markov model/State space model

$$\begin{aligned}x_0 &\sim \mathcal{N}(0, 1) \\x_t|x_{t-1} &\sim \mathcal{N}(x_{t-1}, 1), \quad t = 1, \dots, 10 \\y_t|x_t &\sim \mathcal{N}(x_t, 1), \quad t = 1, \dots, 10\end{aligned}$$

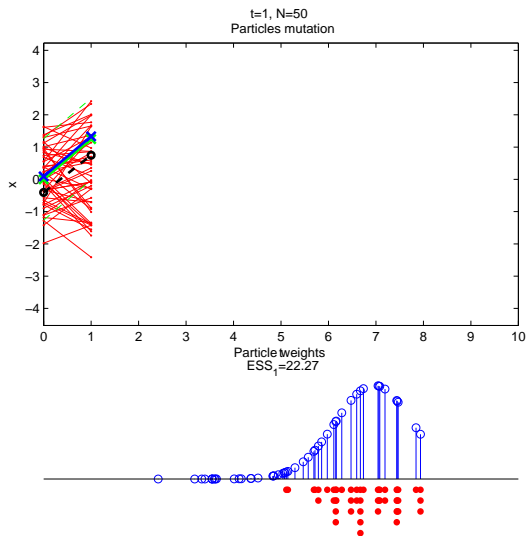


- ▶ Linear Gaussian model
- ▶ Goal: estimate $p(\mathbf{x}_{1:t}|\mathbf{y}_{1:t})$
- ▶ Analytic solution given by Kalman equations

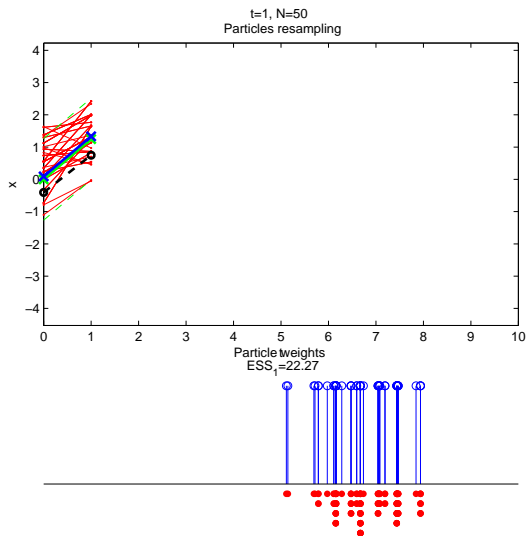
Example: hidden Markov/state space model



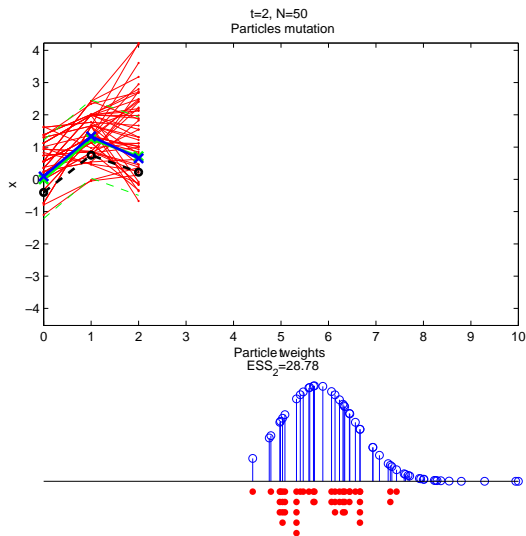
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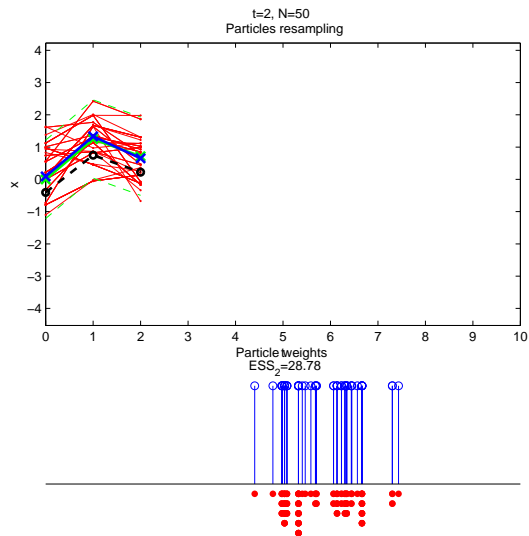
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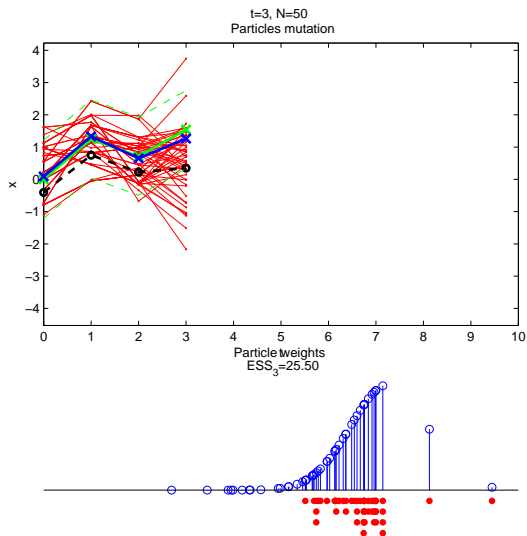
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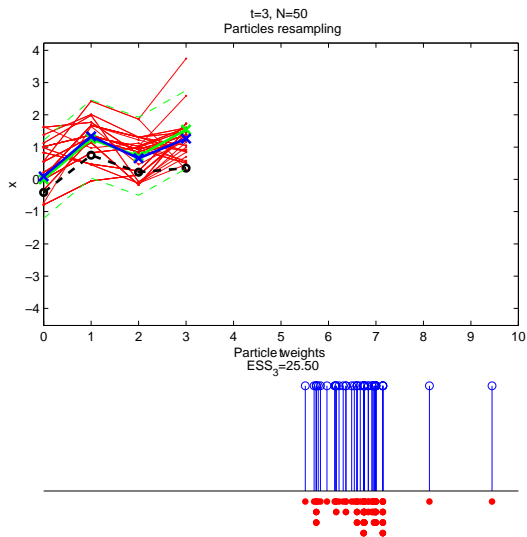
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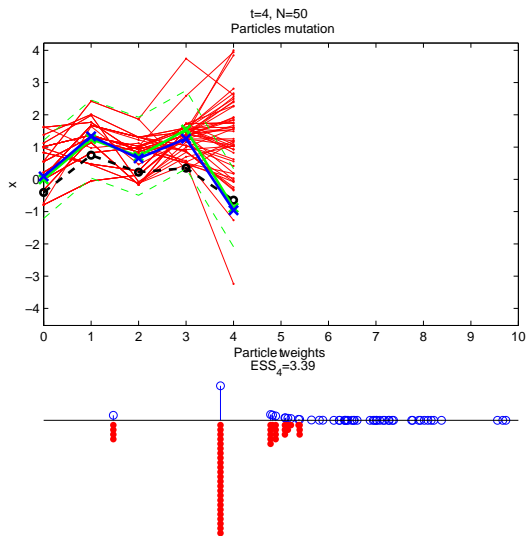
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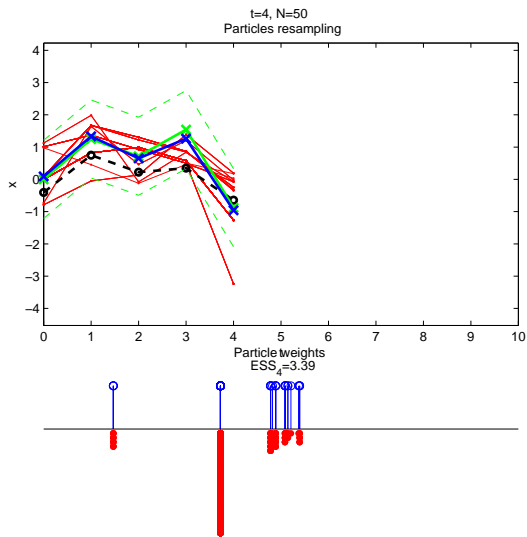
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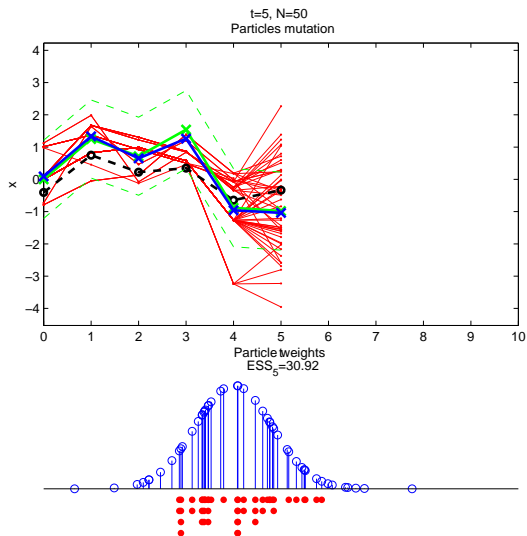
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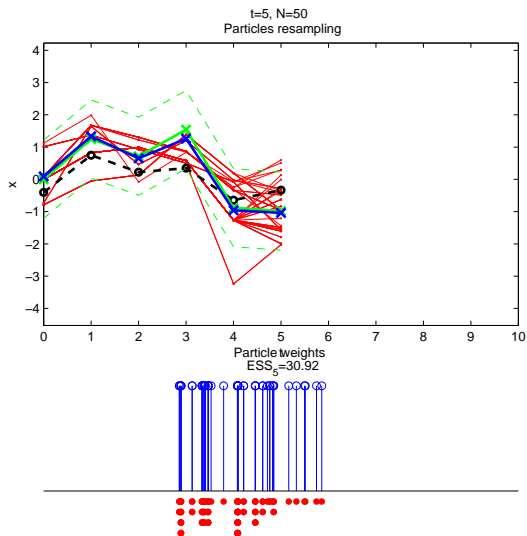
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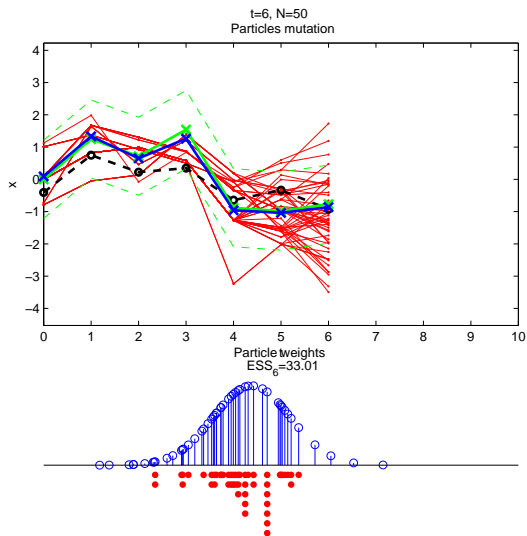
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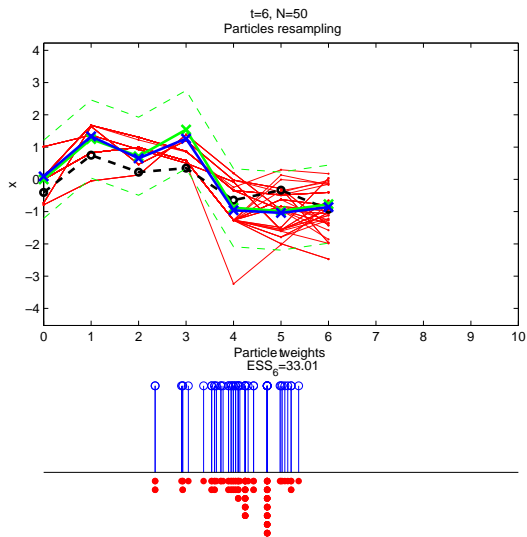
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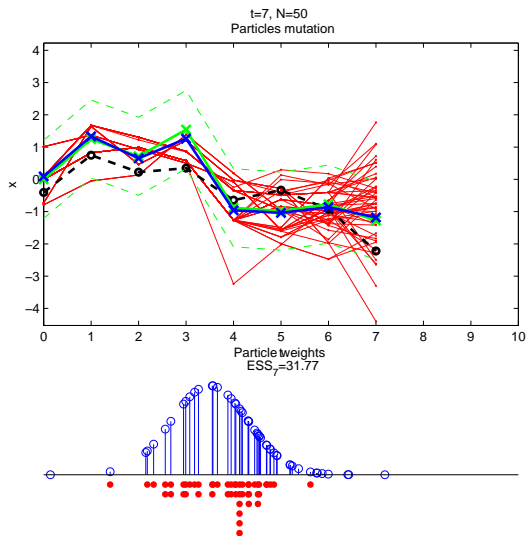
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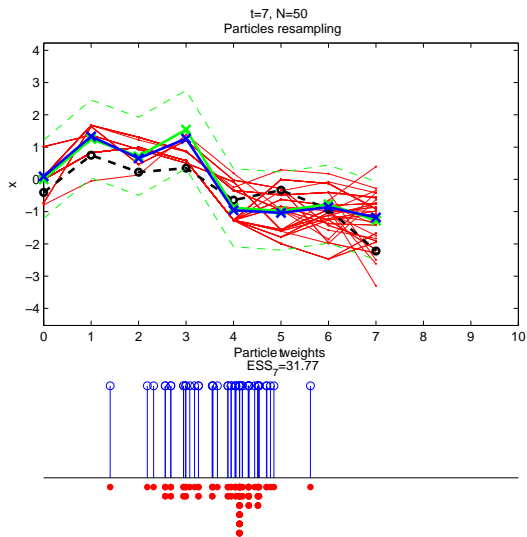
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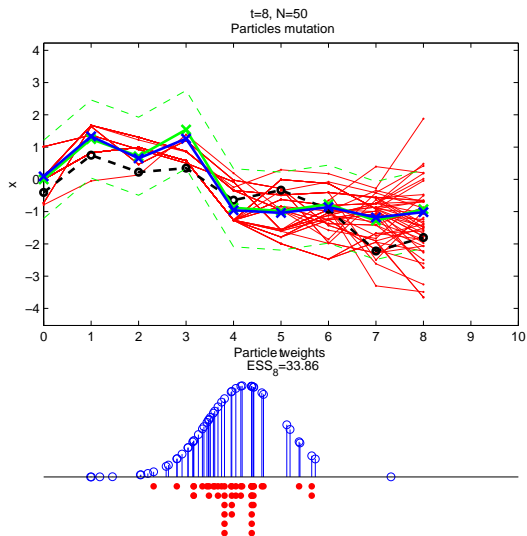
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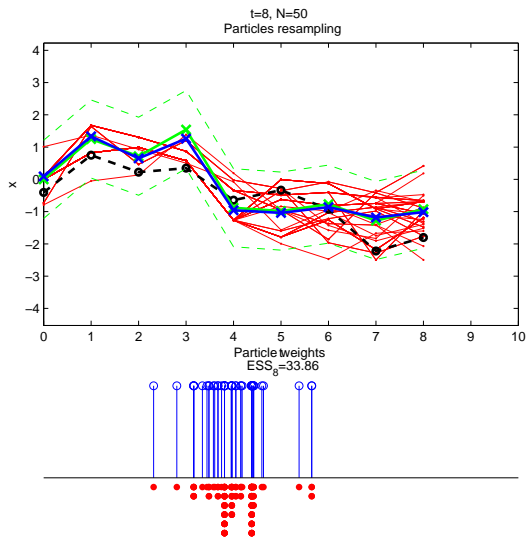
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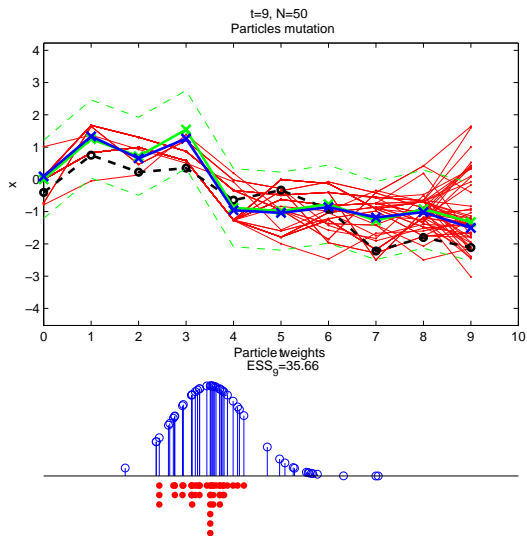
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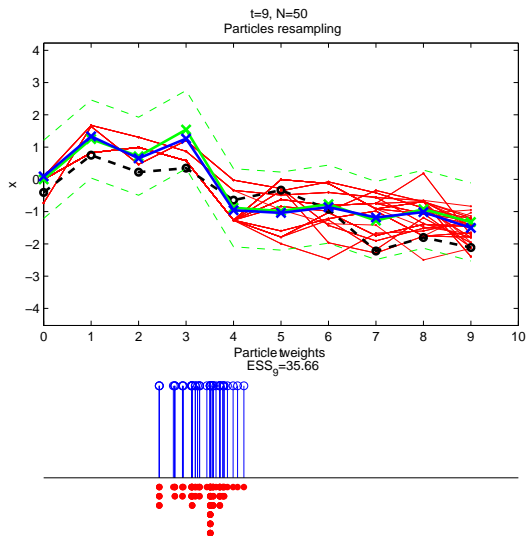
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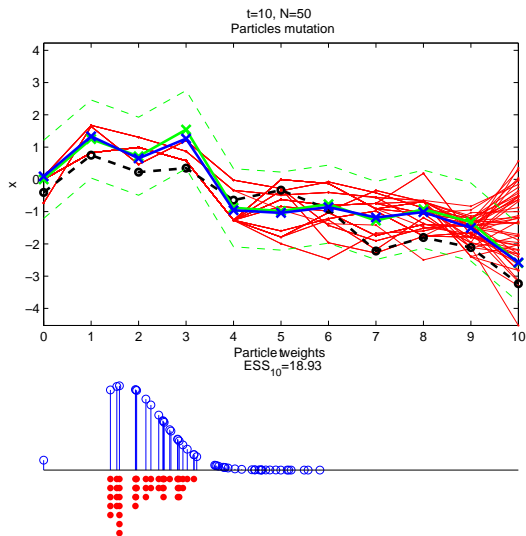
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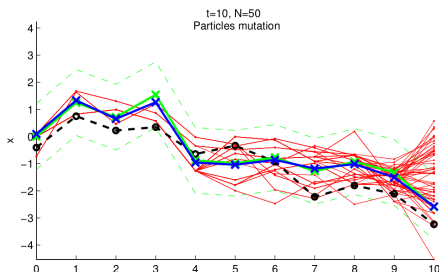
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Limitations and diagnosis of SMC algorithms



For a given $t \leq n$, for each unique value $X_{n,t}'^{(k)}$, $k = 1, \dots, K_{n,t}$, let $W_{n,t}'^{(k)} = \sum_{i|X_t^{(i)}=X_t'^{(k)}} W_n^{(i)}$ be its associated total weight. A measure of the quality of the approximation of the posterior distribution $p(x_{t:n}|y_{1:n})$ is given by the smoothing effective sample size (**SESS**):

$$\text{SESS}_t = \frac{1}{\sum_{k=1}^{K_{n,t}} \left(W_{n,t}'^{(k)} \right)^2} \quad (1)$$

with $1 \leq \text{SESS}_t \leq N$.

Summary

Context

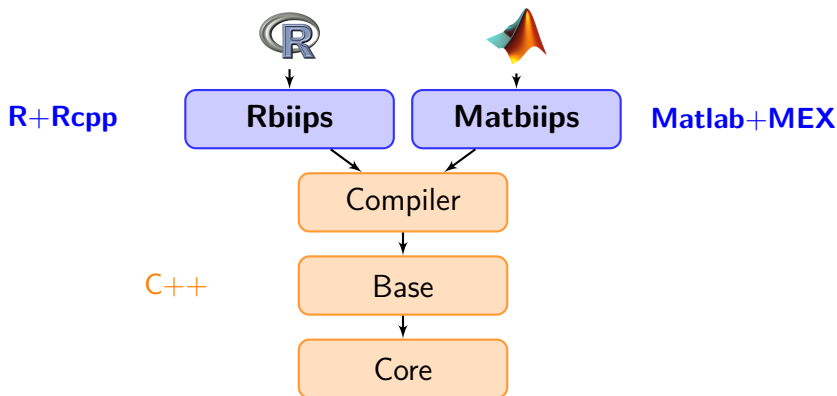
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Particle MCMC

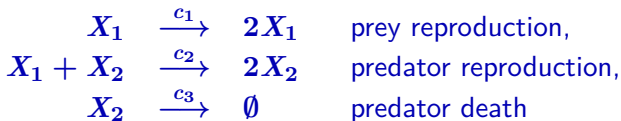
Technical implementation



- ▶ Interfaces: Matlab/Octave, R
- ▶ Multi-platform: Windows, Linux, Mac OSX
- ▶ Free and open source (GPL)

Example: Stochastic kinetic Lotka-Volterra model

- ▶ Evolution of two species $X_1(t)$ (prey) and $X_2(t)$ (predator) at time t
- ▶ Continuous-time Markov jump process described by three reaction equations:



where $c_1 = 0.5$, $c_2 = 0.0025$ and $c_3 = 0.3$.

$$\begin{aligned} \Pr(X_1(t+dt) = x_1(t) + 1, X_2(t+dt) = x_2(t) | x_1(t), x_2(t)) \\ = c_1 x_1(t) dt + o(dt) \end{aligned}$$

$$\begin{aligned} \Pr(X_1(t+dt) = x_1(t) - 1, X_2(t+dt) = x_2(t) + 1 | x_1(t), x_2(t)) \\ = c_2 x_1(t) x_2(t) dt + o(dt) \end{aligned}$$

$$\begin{aligned} \Pr(X_1(t+dt) = x_1(t), X_2(t+dt) = x_2(t) - 1 | x_1(t), x_2(t)) \\ = c_3 x_2(t) dt + o(dt) \end{aligned}$$

[Boys et al., 2008]

Gillespie algorithm

R function to forward simulate from the LV model with Gillespie algorithm

```
lotka_volterra_gillespie <- function(x, c1, c2, c3, dt) {  
  z <- matrix(c(1, -1, 0, 0, 1, -1), nrow=2, byrow=TRUE)  
  t <- 0  
  while (TRUE) {  
    rate <- c(c1*x[1], c2*x[1]*x[2], c3*x[2])  
    sum_rate <- sum(rate);  
    # Sample the next event from an exponential distribution  
    t <- t - log(runif(1))/sum_rate  
    if (t>dt)  
      break  
    # Sample the type of event  
    ind <- which((sum_rate*runif(1)) <= cumsum(rate))[1]  
    x <- x + z[,ind]  
  }  
  return(x)  
}
```

[Gillespie, 1977, Golightly and Gillespie, 2013]

Add a custom sampler to the BUGS language

Rbiips

```
biips_add_distribution(name = 'LV',  
                      n_param = 5,  
                      fun_dim = lotka_volterra_dim,  
                      fun_sample = lotka_volterra_gillespie  
                      )
```


Example: Stochastic kinetic Lotka-Volterra model

- ▶ We observe at some time $t = 1, 2, \dots, t_{\max}$ the number of preys with some additive noise

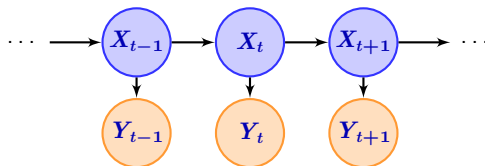
$$Y(t) = X_1(t) + \epsilon(t), \quad \epsilon(t) \sim \mathcal{N}(0, \sigma^2)$$

- ▶ Objective: approximate $\Pr(X_1(t), X_2(t) | Y(1), \dots, Y(t_{\max}))$ at $t = 1, \dots, t_{\max}$.

Example: Stochastic kinetic Lotka-Volterra model

stoch_kinetic_gill.bug

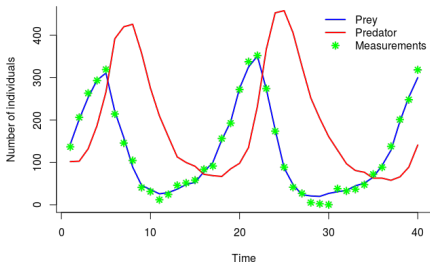
```
model
{
  x[,1] ~ LV(x_init, c[1], c[2], c[3], 1)
  y[1] ~ dnorm(x[1,1], 1/sigma^2)
  for (t in 2:t_max)
  {
    x[,t] ~ LV(x[,t-1], c[1], c[2], c[3], 1)
    y[t] ~ dnorm(x[1,t], 1/sigma^2)
  }
}
```



Model compilation

Rbiips

```
data <- list(t_max = 40, c = c(.5, .0025, .3),  
            x_init = c(100, 100), sigma = 10)  
model <- biips_model(model_file = 'stoch_kinetic_gill.bug',  
                    data = data,  
                    sample_data = TRUE)  
data <- model$data()
```

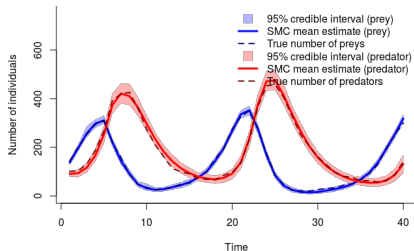


Ground truth and data

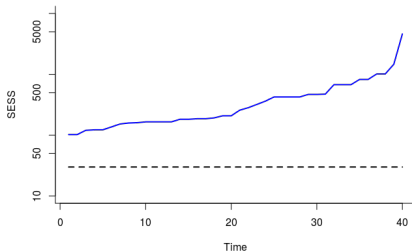
SMC samples

Rbiips

```
out_smc <- biips_smc_samples(model, variable_names = 'x',  
                             n_part = 10000, type = 'fs')  
  
diag_smc <- biips_diagnosis(out_smc)  
summ_smc <- biips_summary(out_smc, probs=c(.025, .975))  
x_s_mean <- summ_smc$x$s$mean  
x_s_quant <- summ_smc$x$s$quant
```



(a) Estimates

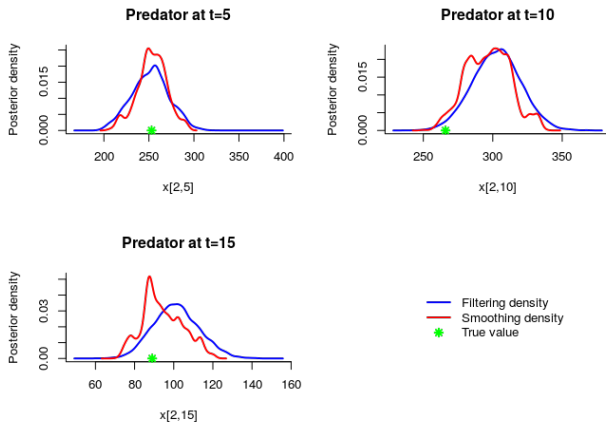


(b) Smoothing effective sample size

Kernel density estimates

Rbiips

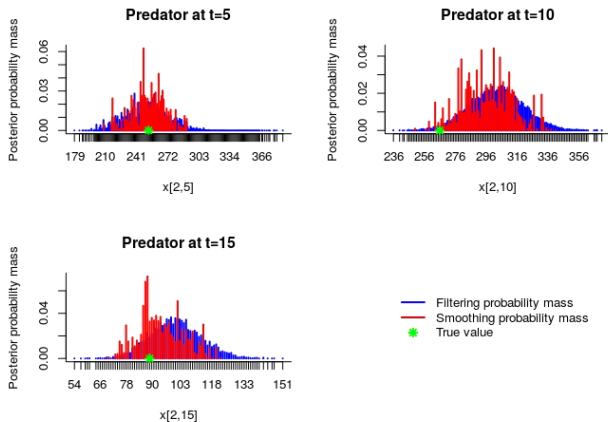
```
kde_smc <- biips_density(out_smc)
```



Probability mass estimates

Rbiips

```
tab_smc <- biips_table(out_smc)
```



Summary

Context

Graphical models and BUGS language

SMC

Biips software

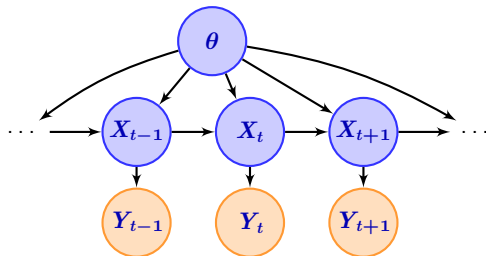
Particle MCMC

Particle MCMC

Recent algorithms that use SMC algorithms within a MCMC algorithm

- ▶ Particle Independent Metropolis-Hastings (PIMH)
- ▶ Particle Marginal Metropolis-Hastings (PMMH)

Static parameter estimation



Due to the successive resamplings, SMC estimations of $p(\theta|y_{1:n})$ might be poor.

The PMMH splits the variables in the graphical model into two sets:

- ▶ a set of variables \mathbf{X} that will be sampled using a SMC algorithm
- ▶ a set $\theta = (\theta_1, \dots, \theta_p)$ sampled with a MH proposal

Standard PMMH algorithm

Set $\hat{Z}(0) = \mathbf{0}$ and initialize $\theta(0)$

For $k = 1, \dots, n_{\text{iter}}$,

- ▶ Sample $\theta^* \sim \nu(\cdot | \theta^{(k-1)})$
- ▶ Run a SMC to approximate $p(x_{1:n} | y_{1:n}, \theta^*)$ with output $(X_{1:n}^{*(i)}, W_n^{*(i)})_{i=1, \dots, N}$ and $\hat{Z}^* \approx p(y_{1:n} | \theta^*)$
- ▶ With probability

$$\min \left(1, \frac{\nu(\theta^* | \theta(k-1)) p(\theta^*) \hat{Z}^*}{\nu(\theta(k-1) | \theta^*) p(\theta(k-1)) \hat{Z}(k-1)} \right)$$

set $X_{1:n}(k) = X_{1:n}^{*(\ell)}$, $\theta(k) = \theta^*$ and $\hat{Z}(k) = \hat{Z}^*$, where $\ell \sim \text{Discrete}(W_n^{*(1)}, \dots, W_n^{*(N)})$

- ▶ otherwise, keep previous iteration values

Outputs

- ▶ MCMC samples $(X_{1:n}(k), \theta(k))_{k=1, \dots, n_{\text{iter}}}$

Example: Stochastic kinetic Lotka-Volterra model

stoch_kinetic_gill.bug

```
model
{
  logc [1] ~ dunif (-7, 2)
  logc [2] ~ dunif (-7, 2)
  logc [3] ~ dunif (-7, 2)
  c [1] <- exp(logc [1])
  c [2] <- exp(logc [2])
  c [3] <- exp(logc [3])
  ...
}
```

Run a PMMH algorithm

Rbiips

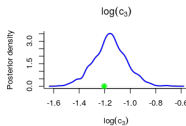
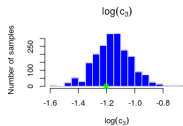
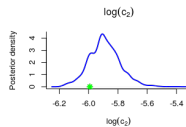
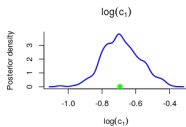
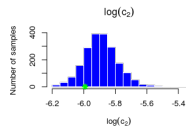
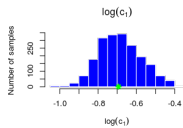
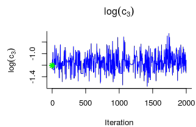
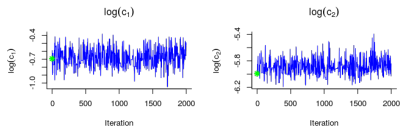
```
# create a pmmh object
obj_pmmh = biips_pmmh_init(model,
                           param_names = c('logc[1]',
                                             'logc[2]',
                                             'logc[3]'),
                           inits = list(-1, -5, -1),
                           latent_names = 'x')

# adaptation and burn-in iterations
biips_pmmh_update(obj_pmmh, n_iter = 2000, n_part = 100)

# samples
out_pmmh = biips_pmmh_samples(obj_pmmh, n_iter = 20000,
                              n_part = 100, thin = 10)

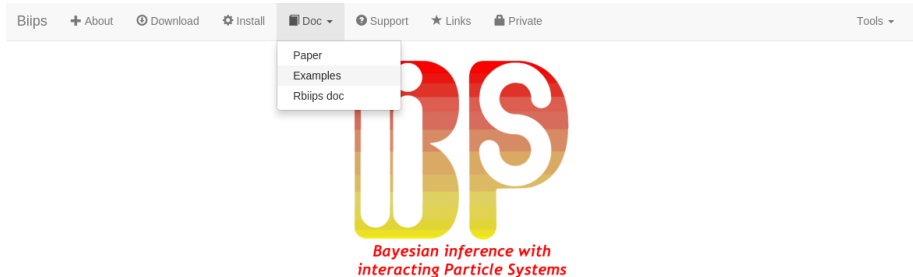
summ_pmmh = biips_summary(out_pmmh, probs = c(.025, .975))
kde_pmmh = biips_density(out_pmmh)
```

Posterior samples



Conclusion

- ▶ BUGS language compatible
- ▶ Extensibility: custom functions/samplers
- ▶ Black-box SMC inference engine
- ▶ Interfaces with popular software: Matlab/Octave, R
- ▶ Post-processing tools
- ▶ And more: backward smoothing algorithm, particle independent Metropolis-Hastings algorithm, sensitivity analysis, some optimal/conditional samplers (Gaussian-Gaussian, beta-Bernoulli, finite discrete)



What is Biips?

Biips is a general software for **Bayesian inference with interacting particle systems**, a.k.a. sequential Monte Carlo (SMC) methods. It aims at popularizing the use of these methods to non-statistician researchers and students, thanks to its automated "black box" inference engine.

It borrows from the [QBUGS/QJAGS](#) software, widely used in Bayesian statistics, the statistical modeling with graphical models and the language associated with their descriptions.

Features

- BUGS language compatible
- SMC techniques for filtering and smoothing
- Static parameter estimation using particle MCMC
- Core developped in C++
- R, Matlab/Octave interfaces
- Easy language extensions with custom R and Matlab functions
- Multi-platform: Linux, Windows, Mac
- Free and open source ([GPL](#))

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THANK YOU



<http://alea.bordeaux.inria.fr/biips>